

# EEE 598C: Statistical Pattern Recognition

## Lecture Note 6a: More Non-parametric Estimation

Darryl Morrell

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### 1 $K_n$ Nearest Neighbor Estimation

Our estimator of the density at point  $\mathbf{x}$  using a training set of size  $n$  is

$$\hat{f}_n(\mathbf{x}) = \frac{k_n}{nV_n}$$

where  $k_n$  is the number of training vectors in a region  $\mathcal{R}_n$  that contains  $\mathbf{x}$  and has volume  $V_n$ . The Parzen window approach was to make  $V_n$  a given function of  $n$ , and then count the number of training vectors in  $\mathcal{R}_n$ . Another approach, called the  $k_n$  nearest neighbor estimate, is to make  $k_n$  a given function of  $n$ , and let  $\mathcal{R}_n$  (and consequently  $V_n$ ) grow until  $k_n$  training vectors are contained in  $\mathcal{R}_n$ .

### 2 Series Expansions

Another method of estimating a density from training data is to find a series expansion of the Parzen window  $\varphi\left(\frac{\mathbf{x}-\mathbf{x}_i}{h_n}\right)$ :

$$\varphi\left(\frac{\mathbf{x}-\mathbf{x}_i}{h_n}\right) = \sum_{j=1}^m a_j \psi_j(\mathbf{x}) \beta_j(\mathbf{x}_i)$$

$a_j$  are the series expansion coefficients.  $\psi_j(\mathbf{x})$  and  $\beta_j(\mathbf{x}_i)$  can be obtained, for example, using Taylor series expansions or other polynomial approximations. With this expansion for the Parzen window, the estimator of the density becomes

$$\begin{aligned} \hat{f}_n(\mathbf{x}) &= \frac{1}{nV_n} \sum_{i=1}^n \varphi\left(\frac{\mathbf{x}-\mathbf{x}_i}{h_n}\right) \\ &= \frac{1}{nV_n} \sum_{i=1}^n \sum_{j=1}^m a_j \psi_j(\mathbf{x}) \beta_j(\mathbf{x}_i) \end{aligned}$$

$$\begin{aligned}
&= \sum_{j=1}^m \psi_j(\mathbf{x}) \left[ \frac{a_j}{nV_n} \sum_{i=1}^n \beta_j(\mathbf{x}_i) \right] \\
&= \sum_{j=1}^m b_j \psi_j(\mathbf{x})
\end{aligned}$$

where

$$b_j = \frac{a_j}{nV_n} \sum_{i=1}^n \beta_j(\mathbf{x}_i)$$

So the information in the training set is summarized in the coefficients  $b_j$ .

### 3 Estimating Posterior Probabilities

Given a set of  $n$  training data  $\mathcal{X}_n$ , how might we find an estimate of  $P(\omega_i|\mathbf{x})$ ? One approach would be to first estimate the joint distribution

$$f(\mathbf{x}, \omega_i) = f(\mathbf{x}|\omega_i)P(\omega_i)$$

and then find the conditional distribution

$$P(\omega_i|\mathbf{x}) = \frac{f(\mathbf{x}, \omega_i)}{\sum_{j=1}^c f(\mathbf{x}, \omega_j)}$$

We could estimate this joint distribution by counting the number of training vectors that fall in a region  $\mathcal{R}_n$  that includes  $\mathbf{x}$  and has volume  $V_n$  as follows:

$$\begin{aligned}
P(\mathbf{X} \in \mathcal{R}_n, \omega_i) &\approx V_n f(\mathbf{x}, \omega_i) \\
f(\mathbf{x}, \omega_i) &\approx \frac{P(\mathbf{X} \in \mathcal{R}_n, \omega_i)}{V_n}
\end{aligned}$$

Let  $k_i$  be the number of training vectors of class  $i$  that fall in  $\mathcal{R}_n$ ; an estimate of  $P(\mathbf{X} \in \mathcal{R}_n, \omega_i)$  is

$$\hat{P}(\mathbf{X} \in \mathcal{R}_n, \omega_i) = \frac{k_i}{n}$$

Thus, an estimate of  $f(\mathbf{x}, \omega_i)$  is

$$\hat{f}_n(\mathbf{x}, \omega_i) = \frac{k_i}{nV_n}$$

and an estimate of  $P(\omega_i|\mathbf{x})$  is

$$\begin{aligned}
\hat{P}_n(\omega_i|\mathbf{x}) &= \frac{\hat{f}_n(\mathbf{x}, \omega_i)}{\sum_{j=1}^c \hat{f}_n(\mathbf{x}, \omega_j)} \\
&= \frac{\frac{k_i}{nV_n}}{\sum_{j=1}^c \frac{k_j}{nV_n}} \\
&= \frac{k_i}{\sum_{j=1}^c k_j}
\end{aligned}$$

Note that  $\sum_{j=1}^c k_j$  is the number of training vectors in  $\mathcal{R}$ . Thus, the estimate of  $P(\omega_i|\mathbf{x})$  is the relative frequency of training vectors from state of nature  $\omega_i$  in the region containing  $\mathbf{x}$ .

We can take either a Parzen window like approach, in which the volume  $V_n$  is a fixed function of  $n$  and we count the number of training vectors in this volume, or a  $k_n$  nearest neighbor approach in which the number of training vectors is fixed as a function of  $n$ , and the volume is increased until this number of training vectors is included.

If the estimate  $\hat{P}_n(\omega_i|\mathbf{x})$  is used in a decision rule, the decision rule is to choose  $\omega_i$  if

$$\hat{P}_n(\omega_i|\mathbf{x}) \geq \hat{P}_n(\omega_j|\mathbf{x})$$

Substituting in the above expression for the estimate, we get the following equivalent decision rule: choose  $\omega_i$  if  $k_i \geq k_j$ ; in other words, if there are more training vectors in  $\mathcal{R}_n$  from class  $\omega_i$  than any other class, choose  $\omega_i$  as the true state of nature.

An approximation to this rule is the *nearest neighbor decision rule*: find the training vector that is nearest  $\mathbf{x}$ , and choose the class of this training vector as the class for  $\mathbf{x}$ . It can be shown that this nearest neighbor decision rule has the follow upper bounds on the probability of error:

$$P_{E_{NN}} \leq 2P_{E_{Bayes}}$$

$$P_{E_{NN}} \leq P_{E_{Bayes}} \left( 2 - \frac{c}{c-1} P_{E_{Bayes}} \right)$$

## 4 Dimensionality Reduction

Non-parametric estimation methods become quite difficult to use with feature vectors of high dimensionality, particularly when a large number of training vectors is not available. One method of reducing the dimensionality of the space is to project the training vectors onto a line chosen to maximize the difference between classes.

For this development, we assume that there are two sets of training data  $\mathcal{X}_1$  and  $\mathcal{X}_2$ , with  $n_1$  and  $n_2$  training vectors. We will project the vectors in these two sets onto a line represented by the vector  $\mathbf{w}$ ; the inner product  $\mathbf{w}^T \mathbf{x}$  gives the position of  $\mathbf{x}$  projected onto the line. By projecting the vectors in  $\mathcal{X}_1$  and  $\mathcal{X}_2$  onto  $\mathbf{w}$ , we obtain two sets of scalars  $\mathcal{Y}_1$  and  $\mathcal{Y}_2$ . Our goal is to choose  $\mathbf{w}$  so as to maximize the distance between these two sets.

One measure of the distance between the sets  $\mathcal{Y}_1$  and  $\mathcal{Y}_2$  is the distance between their sample means. We denote the sample mean of  $\mathcal{Y}_i$  as

$$\begin{aligned} \tilde{m}_i &= \frac{1}{n_i} \sum_{y \in \mathcal{Y}_i} y \\ &= \frac{1}{n_i} \sum_{\mathbf{x} \in \mathcal{X}_i} \mathbf{w}^T \mathbf{x} \\ &= \mathbf{w}^T \left( \frac{1}{n_i} \sum_{\mathbf{x} \in \mathcal{X}_i} \mathbf{x} \right) \\ &= \mathbf{w}^T \mathbf{m}_i \end{aligned}$$

where  $\mathbf{m}_i$  is the sample mean of  $\mathcal{X}_i$ . The distance between sample means is

$$|\tilde{m}_1 - \tilde{m}_2| = \left| \mathbf{w}^T (\tilde{\mathbf{m}}_1 - \tilde{\mathbf{m}}_2) \right|$$

Note that we can make this distance arbitrarily large by scaling  $\mathbf{w}$ . Thus, to find a measure of distance that is invariant to the magnitude of  $\mathbf{w}$ , we also define the *scatter*  $\tilde{s}_i^2$  of each set  $\mathcal{Y}_i$ :

$$\tilde{s}_i^2 = \sum_{y \in \mathcal{Y}_i} (y - \tilde{m}_i)^2$$

Note that the scatter is the unnormalized sample variance of  $\mathcal{Y}_i$ .

We define a criterion function  $J(\mathbf{w})$  as the following:

$$J(\mathbf{w}) = \frac{|\tilde{m}_1 - \tilde{m}_2|^2}{\tilde{s}_1^2 + \tilde{s}_2^2}$$

This is a measure of the separation of  $\mathcal{Y}_1$  and  $\mathcal{Y}_2$  that is invariant to the magnitude of  $\mathbf{w}$ . We wish to find  $\mathbf{w}$  to maximize  $J(\mathbf{w})$ ; to do so we define the following scatter matrices:

$$S_i = \sum_{\mathbf{x} \in \mathcal{X}_i} (\mathbf{x} - \mathbf{m}_i) (\mathbf{x} - \mathbf{m}_i)^T$$

$$S_W = S_1 + S_2$$

$$S_B = (\mathbf{m}_1 - \mathbf{m}_2) (\mathbf{m}_1 - \mathbf{m}_2)^T$$

$S_W$  is called the within class scatter, and  $S_B$  is called the between class scatter. With these definitions, we can see that

$$\tilde{s}_1^2 + \tilde{s}_2^2 = \mathbf{w}^T S_1 \mathbf{w} + \mathbf{w}^T S_2 \mathbf{w} = \mathbf{w}^T S_W \mathbf{w}$$

$$(\tilde{m}_1 - \tilde{m}_2)^2 = \mathbf{w}^T (\mathbf{m}_1 - \mathbf{m}_2) (\mathbf{m}_1 - \mathbf{m}_2)^T \mathbf{w} = \mathbf{w}^T S_B \mathbf{w}$$

The criterion function can be written as

$$J(\mathbf{w}) = \frac{\mathbf{w}^T S_W \mathbf{w}}{\mathbf{w}^T S_B \mathbf{w}}$$

This is immediately recognized as a generalized Rayleigh quotient; the  $\mathbf{w}$  that maximizes  $J(\mathbf{w})$  must satisfy the following generalized eigenvalue problem:

$$S_B \mathbf{w} = \lambda S_W \mathbf{w}$$

If  $S_W$  has an inverse, this problem can be converted to a conventional eigenvalue problem

$$S_W^{-1} S_B \mathbf{w} = \lambda \mathbf{w}$$

Rather than solve this problem directly, we observe that

$$S_B \mathbf{w} = (\mathbf{m}_1 - \mathbf{m}_2) (\mathbf{m}_1 - \mathbf{m}_2)^T \mathbf{w} = \alpha (\mathbf{m}_1 - \mathbf{m}_2)$$

where  $\alpha = (\mathbf{m}_1 - \mathbf{m}_2)^T \mathbf{w}$  is a scalar. Thus, the minimizing  $\mathbf{w}$  is

$$\mathbf{w} = \frac{\alpha}{\lambda} S_W^{-1} (\mathbf{m}_1 - \mathbf{m}_2)$$

Since the magnitude of  $\mathbf{w}$  does not affect the value of  $J(\mathbf{w})$ , we can ignore the scalar constant  $\frac{\alpha}{\lambda}$  and write

$$\mathbf{w} = S_W^{-1} (\mathbf{m}_1 - \mathbf{m}_2)$$

For the case where we have  $c$  training sets  $\mathcal{X}_1$  through  $\mathcal{X}_c$ , we project the  $d$  dimensional feature vectors onto a  $c - 1$  dimensional space. This projection can be represented as a  $d \times c - 1$  matrix  $W$ :

$$W = \begin{bmatrix} \mathbf{w}_1 & \dots & \mathbf{w}_{c-1} \end{bmatrix}$$

The projection can be written as

$$\mathbf{y} = W^T \mathbf{x}$$

We begin to find the matrix  $W$  by defining the following matrices:

$$S_i = \sum_{\mathbf{x} \in \mathcal{X}_i} (\mathbf{x} - \mathbf{m}_i) (\mathbf{x} - \mathbf{m}_i)^T$$

where  $\mathbf{m}_i$  is the sample mean of  $\mathcal{X}_i$ . We also define a total mean vector  $\mathbf{m}$  and a total scatter matrix  $S_T$  as

$$\mathbf{m} = \frac{1}{n} \sum_{\mathbf{x}} \mathbf{x} = \frac{1}{n} \sum_{i=1}^c n_i \mathbf{m}_i$$

$$S_T = \sum_{\mathbf{x}} (\mathbf{x} - \mathbf{m}) (\mathbf{x} - \mathbf{m})^T$$

We define the between class scatter  $S_B$  as

$$S_B = S_T - S_W = \sum_{i=1}^c n_i (\mathbf{m}_i - \mathbf{m}) (\mathbf{m}_i - \mathbf{m})^T$$

We define the objective function  $J(W)$  as

$$J(W) = \frac{|W^T S_B W|}{|W^T S_W W|}$$

To find the columns  $\mathbf{w}_i$  of the matrix  $W$ , we must solve for the largest eigenvalues of the generalized eigenvalue problem

$$S_B \mathbf{w}_i = \lambda_i S_W \mathbf{w}_i$$

In order to solve for  $\mathbf{w}_i$ , one can solve the following equation for  $\lambda_i$

$$|S_B - \lambda_i S_W| = 0$$

and then solve

$$(S_B - \lambda_i S_W) \mathbf{w}_i = 0$$